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# MODELING DIFFERENCES BETWEEN RESPONSE TIMES OF CORRECT AND INCORRECT RESPONSES

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- The benefit of considering RT
  - provide collateral information for **the estimation of ability**
  - shed further light on **the cognitive processes** that led to the observed response

How to model RT and RA data ?

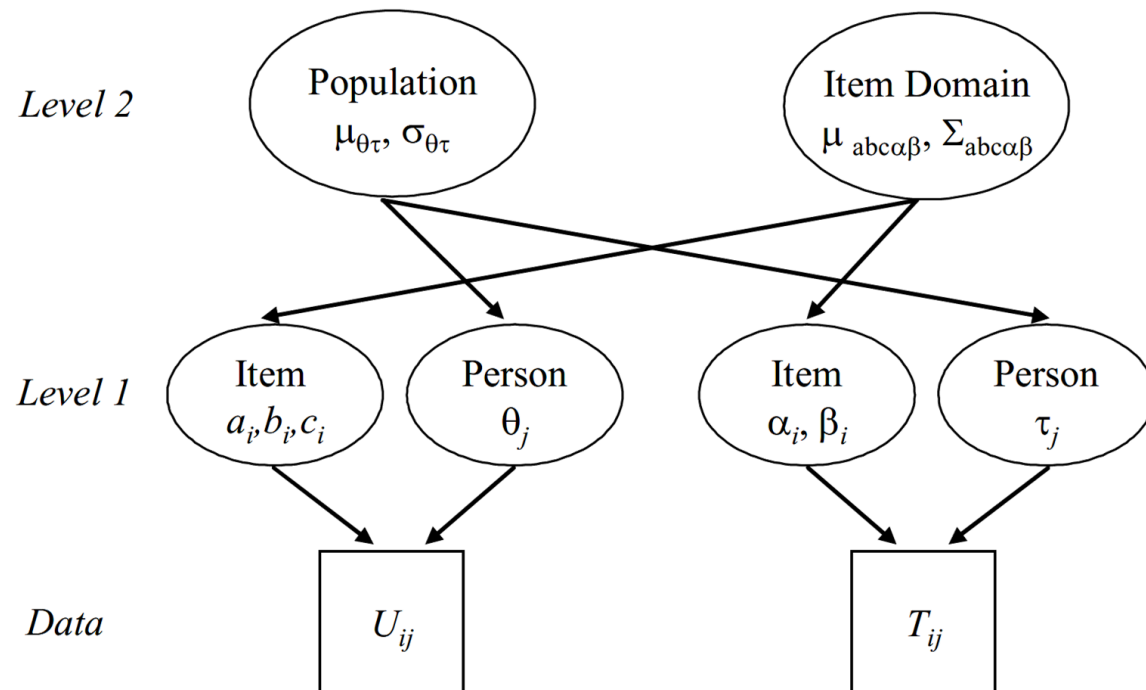
➡  $f(\mathbf{X} = \mathbf{x}, \mathbf{T} = \mathbf{t} \mid \Theta = \theta, \mathbf{H} = \eta)$

- The assumption of independence
  - standard IRT models: given the ability → the RA on different items
  - the lognormal model: given the speed → the RT of different items

- When considering both RA and RT data

How for each item RT & RA are related ?

➔  $f(x_i, t_i | \theta, \eta) = f(x_i | \theta, \eta) f(t_i | \theta, \eta)$



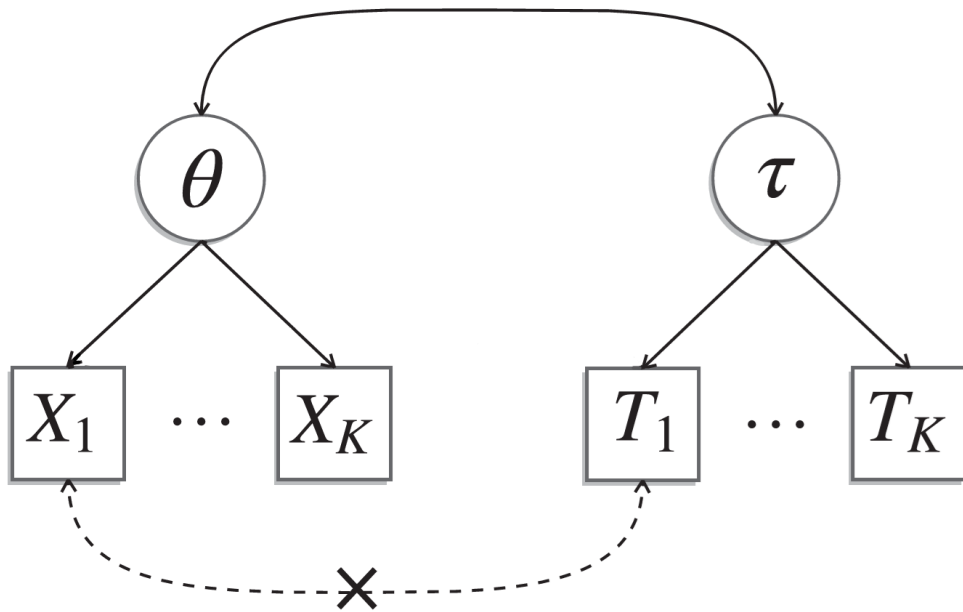
– RA model:

$$\Phi(\alpha_i\theta + \beta_i)^{x_i} (1 - \Phi(\alpha_i\theta + \beta_i))^{1-x_i}$$

– RT model:

$$\ln \mathcal{N}(t_i; \xi_i - \lambda_i\eta, \sigma_i^2)$$

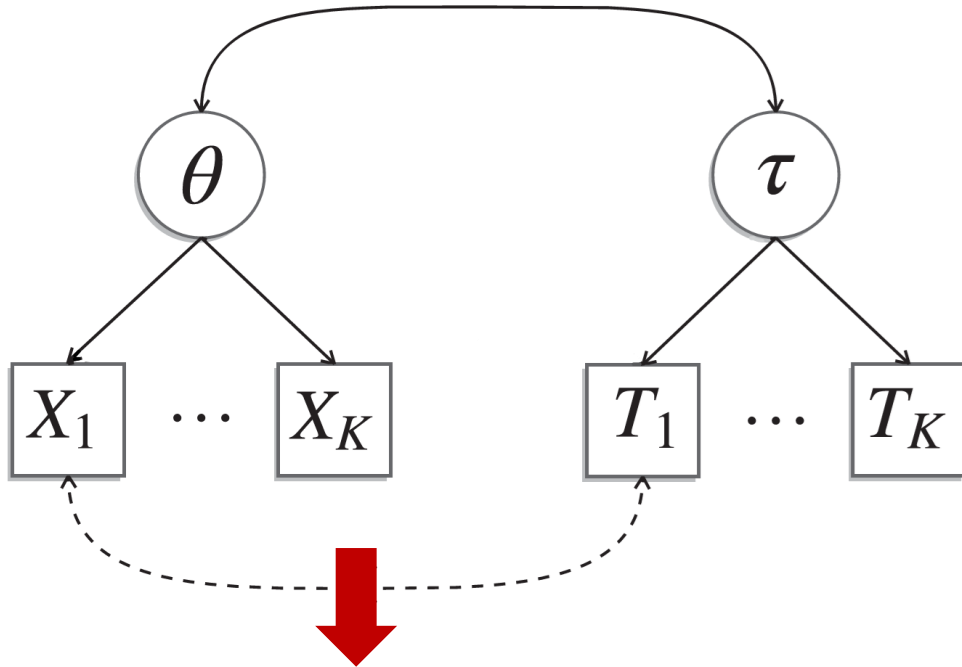
- Residual associations between RA and RT



- speed up during the test
- a temporary lapse in concentration
- differential item functioning
- change problem solving strategies

How to extend the hierarchical modeling framework for RT and RA to allow for conditional dependence (CD) between the outcome variables?

- Conditional Dependence (CD)



1. a bivariate distribution with a nonzero dependence parameter;
2. a marginal distribution of RT and a conditional distribution of RA given RT;
3. the marginal distribution of RA and the conditional distribution of RT given RA.

1. A bivariate distribution with a nonzero dependence parameter

$$f(x_i^*, t_i^* | \theta, \eta) = \mathcal{N}_2 \left( \begin{bmatrix} \alpha_i \theta + \beta_i \\ \xi_i - \lambda_i \eta \end{bmatrix}, \begin{bmatrix} 1 & \rho_i \sigma_i \\ \rho_i \sigma_i & \sigma_i^2 \end{bmatrix} \right)$$

$x_i = \mathcal{I}(x_i^* > 0)$     log-RT
varies across items

$$E(z_g | \theta; \beta_{0g}, \beta_{1g}) = \beta_{0g} + \beta_{1g} \theta \quad P(x_g = 1 | \theta; \beta_{0g}, \beta_{1g}) = \int_0^\infty f(z_g | \theta; \beta_{0g}, \beta_{1g}) dz_g = \Phi(\beta_{0g} + \beta_{1g} \theta)$$

2. A marginal distribution of RT and a conditional distribution of RA given RT

$$f(x_i, t_i | \theta, \eta) = f(t_i | \theta, \eta) f(x_i | t_i, \theta, \eta)$$

$f(x_i | t_i, \theta, \eta) = \Psi \left( \alpha_i \theta + \beta_{i0} + \beta_{i1} \frac{\ln t_i - (\xi_i - \eta)}{\sigma_i}; x_i \right)$ 
varies across items
across person

$$\Psi(\cdot; x_i) = \Phi(\cdot)^{x_i} (1 - \Phi(\cdot))^{1-x_i}$$

3. The marginal distribution of RA and the conditional distribution of RT given RA

$$f(x_i, t_i | \theta, \eta) = f(x_i | \theta, \eta) f(t_i | x_i, \theta, \eta)$$

van der Linden and Glas (2010):

separate time intensity parameters **for the correct and incorrect** responses

- ➔ Why is it important to consider correct responses separately from incorrect responses?
- ➔ What are the benefits of doing so?

- Correct and incorrect responses are likely the result of different response processes
  - **a correct response:**  
successfully following the intended solution strategy
  - **an incorrect response:**  
following the intended solution strategy unsuccessfully  
following a different solution strategy than the one intended  
giving up on the item after trying one's best  
failing to attempt to solve the item (e.g., skipping)



- **Different residual variances**

- RTs of correct responses: show more structural patterns
- incorrect responses: may have larger residual variances

- **Different factor loadings**

- RTs of correct responses: are more strongly related to the speed

- **Different time intensities**

- same ability and speed levels:  $RT_{\text{correct}} > RT_{\text{incorrect}}$  or  $RT_{\text{correct}} < RT_{\text{incorrect}}$

- **Different speed latent variables**

- facing with difficult items: long time or little time

- Empirical support
  - Semmes, Davidson and Close (2011)  
correlations between ability and median RT:  
no correlation for correct RTs & positive correlation for incorrect RTs
  - van der Maas and Wagenmakers (2005)  
ability is negatively correlated with the average correct RTs  
not correlated with the average incorrect RTs



**Purpose:** propose a modeling framework in line with the third approach  
model parameters are allowed to **differ depending on the RA**

- The full model

$$f(x_i, t_i | \theta, \eta) = f(x_i | \theta, \eta) f(t_i | x_i, \theta, \eta)$$

- RA model:

$$\Phi(\alpha_i \theta + \beta_i)^{x_i} (1 - \Phi(\alpha_i \theta + \beta_i))^{1-x_i}$$

- RT model:

$$\ln \mathcal{N}(t_i; \xi_i - \lambda_i \eta, \sigma_i^2)$$

How exactly the dependence of  $t_i$  on  $x_i$  is specified?

$$\ln \mathcal{N}(t_i; \xi_{ix_i} - \lambda_{ix_i} \eta_{x_i}, \sigma_{ix_i}^2) = f(t_i | x_i, \theta, \eta)$$

only one speed (two-dim):  $\eta_0 = \eta_1$

two speed (three-dim):  $\eta_0 \neq \eta_1$

		$\sigma_{i0}^2 = \sigma_{i1}^2$	$\sigma_{i0}^2 \neq \sigma_{i1}^2$
$\xi_{i0} = \xi_{i1}$	$\lambda_{i0} = \lambda_{i1}$	the standard HM $\rightarrow \mathcal{M}_1$	$\mathcal{M}_{3b}$
	$\lambda_{i0} \neq \lambda_{i1}$	-	-
$\xi_{i0} \neq \xi_{i1}$	$\lambda_{i0} = \lambda_{i1}$	van der Linden and Glas (2010) $\rightarrow \mathcal{M}_2$	$\mathcal{M}_{4b}$
	$\lambda_{i0} \neq \lambda_{i1}$	$\mathcal{M}_{3a}$	$\mathcal{M}_{4a}$

- The full model

$$f(x_i, t_i | \boldsymbol{\theta}, \boldsymbol{\eta}) = f(x_i | \boldsymbol{\theta}, \boldsymbol{\eta}) f(t_i | x_i, \boldsymbol{\theta}, \boldsymbol{\eta})$$

- RA model:

$$\Phi(\alpha_i \theta + \beta_i)^{x_i} (1 - \Phi(\alpha_i \theta + \beta_i))^{1-x_i}$$

- RT model:

$$\ln \mathcal{N}(t_i; \xi_i - \lambda_i \eta, \sigma_i^2)$$

$$\ln \mathcal{N}(t_i; \xi_{ix_i} - \lambda_{ix_i} \eta_{x_i}, \sigma_{ix_i}^2) = f(t_i | x_i, \boldsymbol{\theta}, \boldsymbol{\eta})$$

only one speed (two-dim):  $\eta_0 = \eta_1$

two speed (three-dim):  $\eta_0 \neq \eta_1$

- The joint distribution

- $(\boldsymbol{\theta}, \boldsymbol{\eta}) \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

$$\boldsymbol{\mu} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad \boldsymbol{\Sigma} = \begin{bmatrix} 1 & \Sigma_{12} \\ \Sigma_{12} & 1 \end{bmatrix}$$

- $(\boldsymbol{\theta}, \eta_0, \eta_1) \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

$$\boldsymbol{\mu} = \begin{bmatrix} 0 \\ 0 \\ \mu_3 \end{bmatrix} \quad \boldsymbol{\Sigma} = \begin{bmatrix} 1 & \Sigma_{12} & \Sigma_{13} \\ \Sigma_{12} & 1 & \Sigma_{23} \\ \Sigma_{13} & \Sigma_{23} & \Sigma_{33} \end{bmatrix}$$

$\mu_3 \equiv 0$  only when  $\xi_{i0} \neq \xi_{i1}$

$\Sigma_{33} \equiv 1$  only when  $\lambda_{i0} \neq \lambda_{i1}$

- Estimation
  - estimated by sampling from **the joint posterior distribution** of the model parameters
  - ✓ point estimate: averages of the sampled values
  - ✓ 95% credible interval: the 2.5% and 97.5% percentiles of the sampled values
  - (Posterior)~(Likelihood) (Prior)

- Likelihood: for the two-dimensional model

$$\begin{aligned}
 & f(\mathbf{x}, \mathbf{t} \mid \boldsymbol{\alpha}, \boldsymbol{\beta}, \boldsymbol{\xi}, \boldsymbol{\lambda}, \sigma^2, \boldsymbol{\mu}, \boldsymbol{\Sigma}) \\
 &= \prod_{p=1}^N \int \int \prod_{i=1}^K \Psi(\alpha_i \theta + \beta_i; x_{pi}) \ln \mathcal{N}(t_{pi}; \xi_{ix_i} - \lambda_{ix_i} \eta_{x_i}, \sigma_{ix_i}^2) \mathcal{N}_2(\theta, \eta; \boldsymbol{\mu}, \boldsymbol{\Sigma}) d\theta d\eta
 \end{aligned}$$

- Likelihood: for the two-dimensional model

$$\begin{aligned}
 & f(\mathbf{x}, \mathbf{t} \mid \boldsymbol{\alpha}, \boldsymbol{\beta}, \boldsymbol{\xi}, \boldsymbol{\lambda}, \sigma^2, \boldsymbol{\mu}, \boldsymbol{\Sigma}) \\
 &= \prod_{p=1}^N \int \int \int \prod_{i=1}^K \Psi(\alpha_i \theta + \beta_i; x_{pi}) \ln \mathcal{N}(t_{pi}; \xi_{ix_i} - \lambda_{ix_i} \eta_{x_i}, \sigma_{ix_i}^2) \mathcal{N}_3(\theta, \eta_0, \eta_1; \boldsymbol{\mu}, \boldsymbol{\Sigma}) d\theta d\eta_0 d\eta_1
 \end{aligned}$$

- Prior: for the item parameters
  - independent semi-conjugate low-informative priors

$$f(\boldsymbol{\alpha}, \boldsymbol{\beta}, \boldsymbol{\xi}, \boldsymbol{\lambda}, \sigma^2)$$

$$= \prod_{i=1}^K \mathcal{N}(\alpha_i; 0, 100^2) \mathcal{N}(\beta_i; 0, 100^2) \prod_{k=\{0,1\}} \mathcal{N}(\xi_{ik}; 0, 100^2) \mathcal{N}(\lambda_{ik}; 0, 100^2) \mathcal{IG}(\sigma_{ik}^2; 0.001, 0.001)$$

- Prior: for the person parameters

 Gibbs Sampler

- $(\theta, \eta_0, \eta_1) \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$



- ✓ the mean vector and the covariance matrix are (partially) constrained
- ✓ sample them freely but for each sample from the posterior rescale all the parameters


- Model Selection


- Akaike information criterion (AIC)

$$\text{AIC} = -2\ln(\text{likelihood function}) + 2(\text{numbers of parameters})$$

- Bayesian information criterion (BIC)

$$\text{BIC} = -2\ln(\text{likelihood function}) + \ln(n)(\text{numbers of parameters})$$

 Bayesian estimation procedure  
at the posterior mean of the parameters

  
mAIC & mBIC



- The values of log-likelihood

– two-dimensional models

$$\ln \mathcal{L}_{2\text{dim}}(\hat{\alpha}, \hat{\beta}, \hat{\xi}, \hat{\lambda}, \hat{\sigma}^2, \hat{\Sigma}_{12}; \mathbf{x}, \mathbf{t})$$

$$\approx \sum_{p=1}^N \sum_{g=1}^{10} \sum_{h=1}^{10} \frac{w_g}{\sqrt{\pi}} \frac{w_h}{\sqrt{\pi}} \prod_{i=1}^K \Psi(\hat{\alpha}_i v_{ghj1} + \hat{\beta}_i; x_{pi}) \ln \mathcal{N}(t_{pi}; \hat{\xi}_{ix_i} - \hat{\lambda}_{ix_i} v_{ghj2}, \hat{\sigma}_{ix_i}^2)$$

$$\begin{aligned} v_{ghj1} &= \sqrt{2}y_g, \\ v_{ghj2} &= \sqrt{2(1 - \hat{\Sigma}_{12}^2)}y_h + \hat{\Sigma}_{12}v_{ghj1}, \\ v_{ghj3} &= \sqrt{2\left(\hat{\Sigma}_{33} - \frac{\hat{\Sigma}_{13}^2 + \hat{\Sigma}_{23}^2 - 2\hat{\Sigma}_{12}\hat{\Sigma}_{13}\hat{\Sigma}_{23}}{1 - \hat{\Sigma}_{12}^2}\right)}y_j + \hat{\mu}_3 + \frac{\hat{\Sigma}_{13} - \hat{\Sigma}_{12}\hat{\Sigma}_{23}}{1 - \hat{\Sigma}_{12}^2}v_{ghj1} + \frac{\hat{\Sigma}_{23} - \hat{\Sigma}_{12}\hat{\Sigma}_{13}}{1 - \hat{\Sigma}_{12}^2}v_{ghj2} \end{aligned}$$

– three-dimensional models

$$\ln \mathcal{L}_{3\text{dim}}(\hat{\alpha}, \hat{\beta}, \hat{\xi}, \hat{\lambda}, \hat{\sigma}^2, \hat{\Sigma}_{12}, \hat{\Sigma}_{13}, \hat{\Sigma}_{23}, \hat{\Sigma}_{33}, \hat{\mu}_3; \mathbf{x}, \mathbf{t})$$

$$\approx \sum_{p=1}^N \sum_{g=1}^{10} \sum_{h=1}^{10} \sum_{j=1}^{10} \frac{w_g}{\sqrt{\pi}} \frac{w_h}{\sqrt{\pi}} \frac{w_j}{\sqrt{\pi}} \prod_{i=1}^K \Psi(\hat{\alpha}_i v_{ghj1} + \hat{\beta}_i; x_{pi}) \ln \mathcal{N}(t_{pi}; \hat{\xi}_{ix_i} - \hat{\lambda}_{ix_i} v_{ghj(2+x_i)}, \hat{\sigma}_{ix_i}^2)$$

- The number of parameters

$$\ln \mathcal{L}_{2\text{dim}}(\hat{\alpha}, \hat{\beta}, \hat{\xi}, \hat{\lambda}, \hat{\sigma}^2, \hat{\Sigma}_{12}; \mathbf{x}, \mathbf{t})$$

$$\ln \mathcal{L}_{3\text{dim}}(\hat{\alpha}, \hat{\beta}, \hat{\xi}, \hat{\lambda}, \hat{\sigma}^2, \hat{\Sigma}_{12}, \hat{\Sigma}_{13}, \hat{\Sigma}_{23}, \hat{\Sigma}_{33}, \hat{\mu}_3; \mathbf{x}, \mathbf{t})$$

– on the item side

5K parameters

+ K with  $\xi_{i0} \neq \xi_{i1}$

+ K with  $\lambda_{i0} \neq \lambda_{i1}$

+ K with  $\sigma^2_{i0} \neq \sigma^2_{i1}$

– on the population side

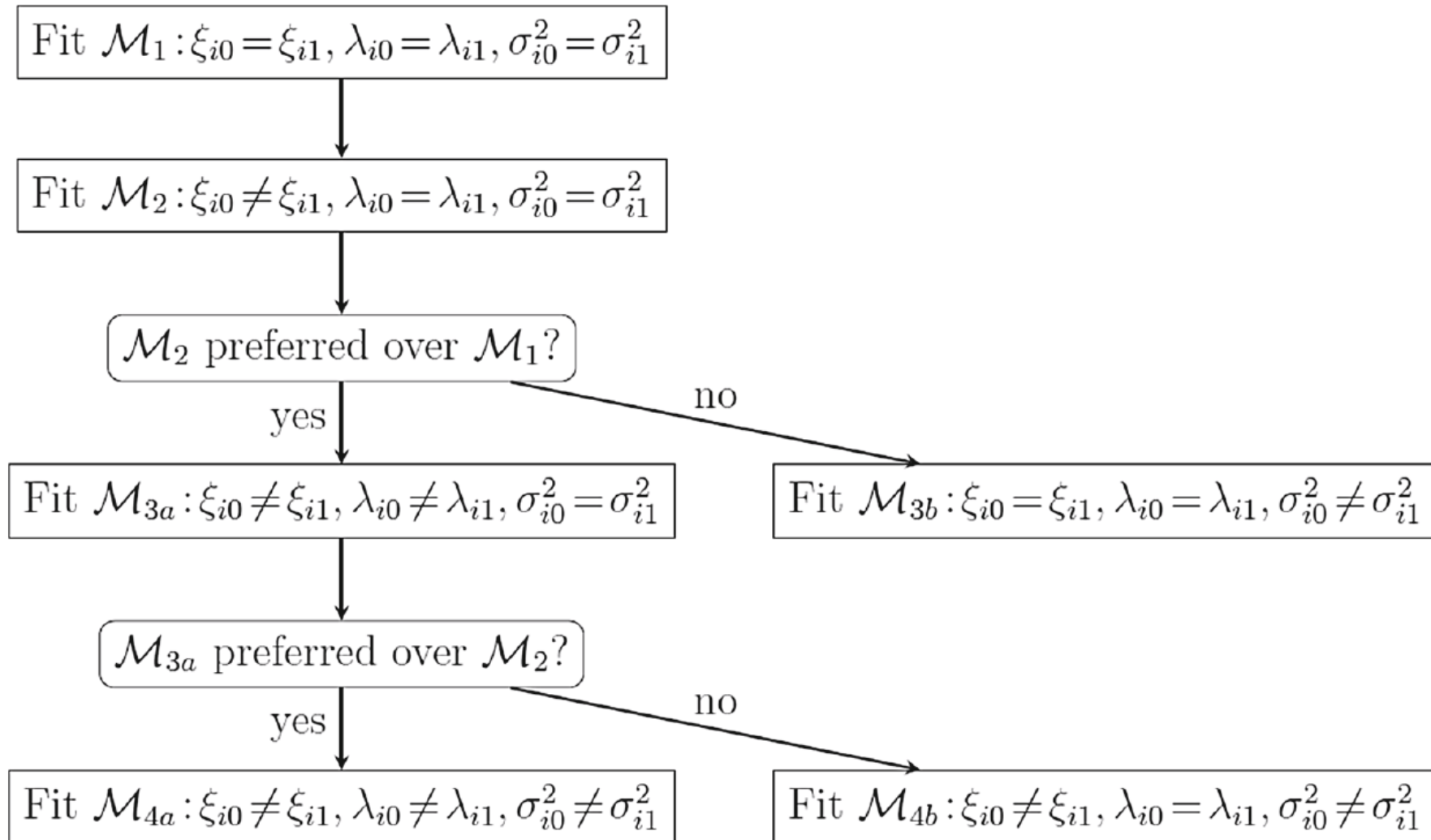
one covariance with  $\eta_0 = \eta_1$

three covariances with  $\eta_0 \neq \eta_1$

one freely estimated mean with  $\eta_0 \neq \eta_1$  &  $\xi_{i0} = \xi_{i1}$

one freely estimated variance with  $\eta_0 \neq \eta_1$  &  $\lambda_{i0} = \lambda_{i1}$

- Stepwise model selection




- Posterior Predictive Check (PPC)

Whether the best two-dimensional model adequately captures the relevant patterns?

– the correlation between persons'  $M_{\text{correct log-RT}}$  and  $M_{\text{incorrect log-RT}}$

1. calculated for the observed data and for  $G$  replicated data sets
2. p-value: the proportion of data sets in which the replicated statistic is **larger than** the observed statistic
3. p-values **close to 1** indicate model misfit: **three-dim** is needed
4. p-value is **below a certain threshold** (e.g., 0.95) indicate model **fit well**

- Simulation Study 1: Parameter Recovery
  - for the two-dimensional and three-dimensional models
- Simulation Study 2: Model Selection
  - generate data under all twelve models
- Empirical Example: PIAAC Problem Solving
- Simulation Study Based on the Empirical Example

- the baseline condition
    - sample size = 1000
    - number of items = 16
    - correlation(s) between speed and ability = 0
    - In the case of the three-dimensional model: correlation between the two speed = 0.7
  - extra conditions
    - twice as large (2000) and twice as small (500)
    - 32 items
    - correlation of 0.5
    - larger correlation between the two speed (0.9)
  - the item parameters
    - $\alpha_i$  values: 0.5 and 1
    - $\{\lambda_{i0}, \lambda_{i1}\} : \{0.3, 0.4\}$  and  $\{0.4, 0.3\}$
    - $\{\xi_{i0}, \xi_{i1}\} : \{4, 4.1\}$  and  $\{4.1, 4\}$
    - $\{\sigma^2_{i0}, \sigma^2_{i1}\} : \{0.3, 0.2\}$  and  $\{0.2, 0.3\}$
    - item intercept parameters  $\beta_i$ : equally spaced between  $-1.5$  and  $1.5$
-  **16 unique item parameter combinations**  
**the same item parameters were used twice with 32 items**

- data sets

- the first five conditions:

- 500 data sets (both the two-dimensional and three-dimensional version of  $\mathcal{M}_{4a}$ )

- the last condition: 500 data sets (the three-dimensional  $\mathcal{M}_{4a}$ )

- RA data:  $\Phi(\alpha_i\theta + \beta_i)^{x_i} (1 - \Phi(\alpha_i\theta + \beta_i))^{1-x_i}$

- RT data:  $\ln \mathcal{N}(t_i; \xi_{ix_i} - \lambda_{ix_i}\eta_{x_i}, \sigma_{ix_i}^2)$

- person parameters

- $\mathcal{N}_2 \left( \mathbf{0}, \begin{bmatrix} 1 & \Sigma_{12} \\ \Sigma_{12} & 1 \end{bmatrix} \right)$  for the conditions with  $\eta_0 = \eta_1$

- $\mathcal{N}_3 \left( \mathbf{0}, \begin{bmatrix} 1 & \Sigma_{12} & \Sigma_{13} \\ \Sigma_{12} & 1 & \Sigma_{23} \\ \Sigma_{13} & \Sigma_{23} & 1 \end{bmatrix} \right)$  for the conditions with  $\eta_0 \neq \eta_1$

- Estimation

- Gibbs Sampler with 6000 iterations

- burn-in: first 1000 iterations

- Evaluation

- the (average) absolute bias

- variance

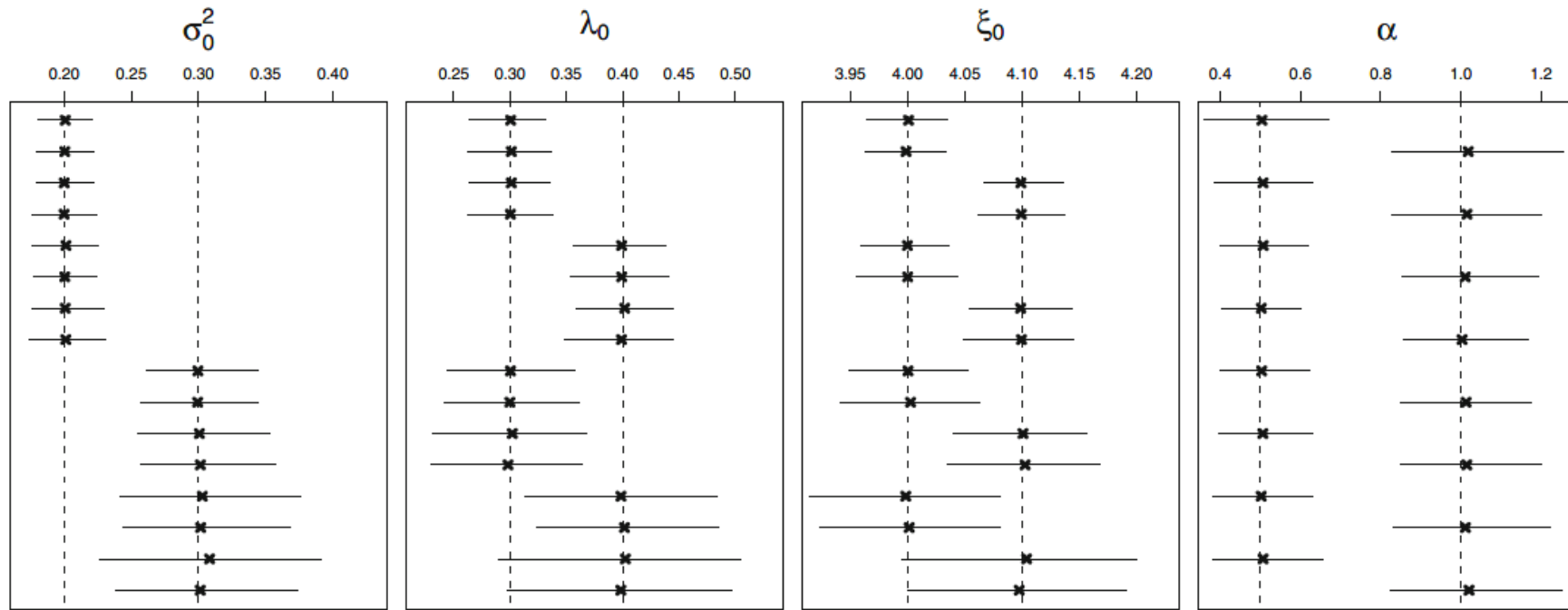
- mean squared error

Condition		$\alpha$	$\beta$	$\xi$	$\lambda$	$\sigma^2$	$\Sigma_{12}$	$\Sigma_{13}$	$\Sigma_{23}$
Bias									
$\eta_0 = \eta_1$	Baseline	0.008	0.007	0.001	0.001	0.002	0.001	-	-
	$N = 500$	0.021	0.014	0.002	0.002	0.003	0.001	-	-
	$N = 2000$	0.005	0.003	0.001	0.000	0.001	0.001	-	-
	$K = 32$	0.008	0.006	0.001	0.001	0.002	0.000	-	-
	$\Sigma_{12} = .5$	0.009	0.007	0.001	0.001	0.002	0.004	-	-
$\eta_0 \neq \eta_1$	Baseline	0.008	0.006	0.001	0.001	0.001	0.001	0.001	0.003
	$N = 500$	0.018	0.014	0.002	0.002	0.004	0.000	0.000	0.007
	$N = 2000$	0.004	0.003	0.001	0.001	0.001	0.002	0.000	0.002
	$K = 32$	0.009	0.006	0.001	0.001	0.002	0.001	0.000	0.002
	$\Sigma_{12} = \Sigma_{13} = .5$	0.010	0.007	0.001	0.001	0.002	0.002	0.001	0.001
	$\Sigma_{23} = .9$	0.011	0.008	0.001	0.001	0.001	0.000	0.002	0.011
Variance									
$\eta_0 = \eta_1$	Baseline	0.006	0.004	0.001	0.001	0.001	0.001	-	-
	$N = 500$	0.013	0.009	0.002	0.002	0.001	0.003	-	-
	$N = 2000$	0.003	0.002	0.000	0.001	0.000	0.001	-	-
	$K = 32$	0.005	0.004	0.001	0.001	0.001	0.001	-	-
	$\Sigma_{12} = .5$	0.007	0.004	0.001	0.001	0.001	0.001	-	-
$\eta_0 \neq \eta_1$	Baseline	0.007	0.004	0.001	0.001	0.001	0.002	0.002	0.001
	$N = 500$	0.014	0.009	0.002	0.002	0.001	0.003	0.004	0.001
	$N = 2000$	0.003	0.002	0.001	0.001	0.000	0.001	0.001	0.000
	$K = 32$	0.005	0.004	0.001	0.001	0.001	0.001	0.001	0.000
	$\Sigma_{12} = \Sigma_{13} = .5$	0.006	0.004	0.001	0.001	0.001	0.001	0.001	0.001
	$\Sigma_{23} = .9$	0.007	0.004	0.001	0.001	0.001	0.002	0.002	0.000
Mean squared error									
$\eta_0 = \eta_1$	Baseline	0.006	0.004	0.001	0.001	0.001	0.001	-	-
	$N = 500$	0.014	0.009	0.002	0.002	0.001	0.003	-	-
	$N = 2000$	0.003	0.002	0.000	0.001	0.000	0.001	-	-
	$K = 32$	0.005	0.004	0.001	0.001	0.001	0.001	-	-
	$\Sigma_{12} = .5$	0.007	0.004	0.001	0.001	0.001	0.001	-	-
$\eta_0 \neq \eta_1$	Baseline	0.007	0.004	0.001	0.001	0.001	0.002	0.002	0.001
	$N = 500$	0.014	0.010	0.002	0.002	0.001	0.003	0.004	0.001
	$N = 2000$	0.003	0.002	0.001	0.001	0.000	0.001	0.001	0.000
	$K = 32$	0.005	0.004	0.001	0.001	0.001	0.001	0.001	0.000
	$\Sigma_{12} = \Sigma_{13} = .5$	0.006	0.004	0.001	0.001	0.001	0.001	0.001	0.001
	$\Sigma_{23} = .9$	0.007	0.005	0.001	0.001	0.001	0.002	0.002	0.000



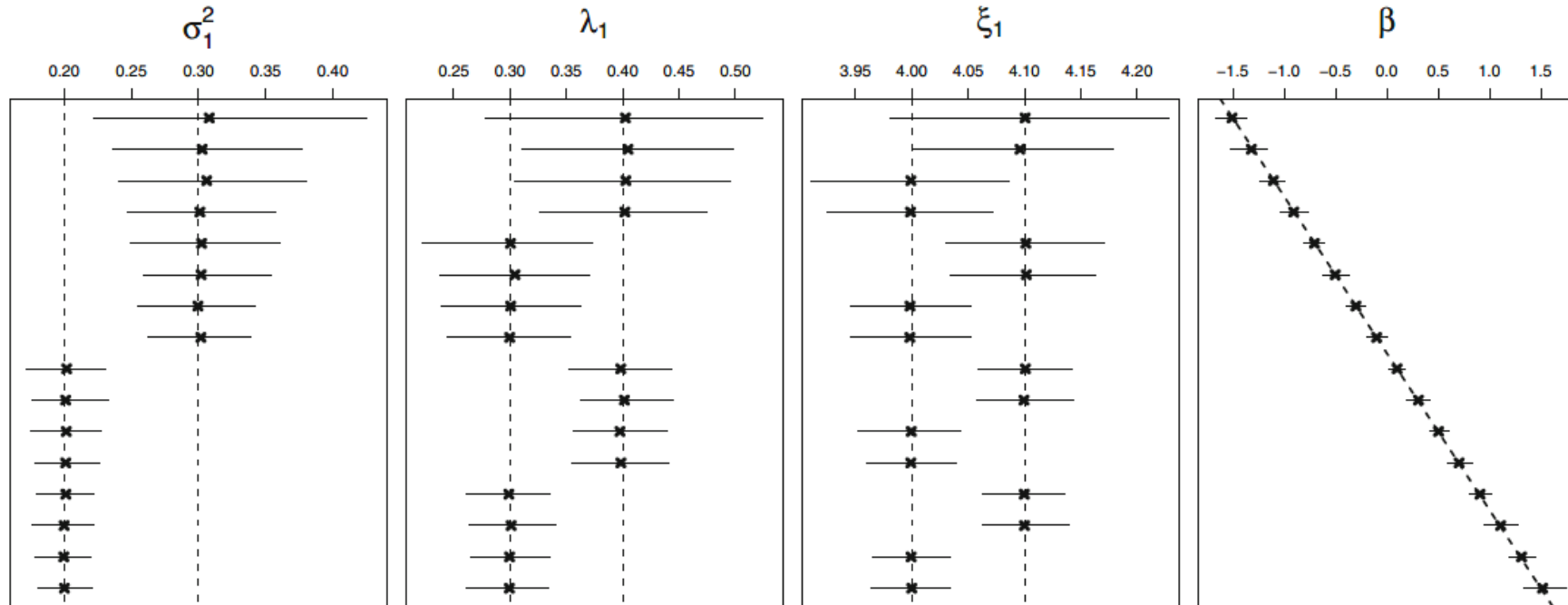
# Results

difficult items →



easy items →

difficult items →



easy items →

- model selection
  - mAIC
  - mBIC
  - mAIC in combination with the posterior predictive check
  - mBIC in combination with the posterior predictive check

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Condition	$N$	$K$	$\Sigma_{23}$
A (baseline)	1000	20	.7
B	1000	<b>10</b>	.7
C	1000	<b>40</b>	.7
D	<b>500</b>	20	.7
E	<b>2000</b>	20	.7
F	1000	20	<b>.9</b>

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Note: For each non-baseline condition the factor that differentiates it from the baseline condition is in bold. Condition F was used only for the three-dimensional models.

- data sets

- each condition 50 data sets were generated under each of the 12 models

- item parameters

- $\alpha_i \sim \mathcal{N}(1, 0.2^2)$

- $\beta_i \sim \mathcal{N}(0, 0.5^2)$

- equal time intensities condition, otherwise:

- $\xi_i \sim \mathcal{N}(4, 0.5^2)$

- $[\xi_{i0} \ \xi_{i1}]^T \sim \mathcal{N}_2 \left( \begin{bmatrix} 4 \\ 4.1 \end{bmatrix}, 0.25 \begin{bmatrix} 1 & 0.7 \\ 0.7 & 1 \end{bmatrix} \right)$

- equal factor loadings condition, otherwise:

- $\lambda_i \sim \mathcal{N}(0.4, 0.1^2)$

- $[\lambda_{i0} \ \lambda_{i1}]^T \sim \mathcal{N}_2 \left( \begin{bmatrix} 0.4 \\ 0.4\sqrt{0.8} \end{bmatrix}, 0.01 \begin{bmatrix} 1 & 0.7 \\ 0.7 & 1 \end{bmatrix} \right)$

- equal residual variances condition, otherwise:

- $\sigma_i^2 \sim \mathcal{U}(0.2, 0.3)$

- $\sigma_{i0}^2 \sim \mathcal{U}(0.2, 0.3)$  and  $\sigma_{i1}^2 \sim \mathcal{U}(0.15, 0.25)$

- person parameters

- for the two-dimensional models:  $\mathcal{N}_2(\mathbf{0}, \mathbf{I}_2)$
- for the three-dimensional models:  $\mathcal{N}_3(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

$$\boldsymbol{\mu} = \mathbf{0} \text{ when } \xi_{i0} \neq \xi_{i1}$$

$$\boldsymbol{\mu} = \begin{bmatrix} 0 \\ 0 \\ 0.1 \end{bmatrix} \text{ when } \xi_{i0} = \xi_{i1}$$

$$\boldsymbol{\Sigma} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & \Sigma_{23}/\sqrt{\Sigma_{33}} \\ 0 & \Sigma_{23}/\sqrt{\Sigma_{33}} & 1 \end{bmatrix} \text{ when } \lambda_{i0} \neq \lambda_{i1}$$

$$\boldsymbol{\Sigma} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & \Sigma_{23}/\sqrt{\Sigma_{33}}\sqrt{0.8} \\ 0 & \Sigma_{23}/\sqrt{\Sigma_{33}}\sqrt{0.8} & 0.8 \end{bmatrix} \text{ when } \lambda_{i0} = \lambda_{i1}$$

- Estimation

- Gibbs Sampler with 6000 iterations
- burn-in: first 1000 iterations
- each second iteration after the burn-in was used

True model		$P$	mAIC						mAIC and PPC					
			Condition						Condition					
			A	B	C	D	E	F	A	B	C	D	E	F
$\eta_0 = \eta_1$	$\mathcal{M}_1$	$5K + 1$	1	41	6	0	0	–	49	48	47	47	48	–
	$\mathcal{M}_2$	$6K + 1$	0	39	4	0	0	–	49	45	48	50	49	–
	$\mathcal{M}_{3a}$	$7K + 1$	5	42	13	0	0	–	49	47	45	49	48	–
	$\mathcal{M}_{3b}$	$6K + 1$	0	28	2	0	0	–	47	46	47	47	49	–
	$\mathcal{M}_{4a}$	$8K + 1$	0	20	1	0	0	–	48	48	49	49	49	–
	$\mathcal{M}_{4b}$	$7K + 1$	0	32	4	0	0	–	49	45	47	50	49	–
$\eta_0 \neq \eta_1$	$\mathcal{M}_1$	$5K + 5$	49	46	49	50	50	50	49	46	49	50	50	42
	$\mathcal{M}_2$	$6K + 4$	50	47	48	49	48	48	50	47	48	49	48	41
	$\mathcal{M}_{3a}$	$7K + 3$	50	50	50	50	49	48	50	49	50	50	49	43
	$\mathcal{M}_{3b}$	$6K + 5$	50	49	49	50	50	50	50	47	49	50	50	47
	$\mathcal{M}_{4a}$	$8K + 3$	49	48	50	50	50	50	50	45	45	50	50	43
	$\mathcal{M}_{4b}$	$7K + 4$	50	46	45	50	50	50	49	47	50	50	50	45

True model		$P$	mBIC						mBIC and PPC					
			Condition						Condition					
			A	B	C	D	E	F	A	B	C	D	E	F
$\eta_0 = \eta_1$	$\mathcal{M}_1$	$5K + 1$	14	50	0	24	0	–	49	50	49	48	50	–
	$\mathcal{M}_2$	$6K + 1$	3	50	0	20	0	–	49	50	48	50	50	–
	$\mathcal{M}_{3a}$	$7K + 1$	3	16	0	1	2	–	24	17	1	29	48	–
	$\mathcal{M}_{3b}$	$6K + 1$	3	33	0	2	0	–	40	33	12	42	49	–
	$\mathcal{M}_{4a}$	$8K + 1$	0	15	0	0	0	–	27	19	3	31	47	–
	$\mathcal{M}_{4b}$	$7K + 1$	0	32	0	2	0	–	42	38	9	49	49	–
$\eta_0 \neq \eta_1$	$\mathcal{M}_1$	$5K + 5$	50	50	50	50	50	50	43	37	48	41	48	45
	$\mathcal{M}_2$	$6K + 4$	50	50	50	50	50	50	39	38	33	38	45	43
	$\mathcal{M}_{3a}$	$7K + 3$	12	10	10	3	41	16	7	8	3	7	20	14
	$\mathcal{M}_{3b}$	$6K + 5$	43	38	50	17	49	43	43	38	17	50	49	41
	$\mathcal{M}_{4a}$	$8K + 3$	19	9	17	0	44	40	21	9	2	15	44	36
	$\mathcal{M}_{4b}$	$7K + 4$	45	42	50	18	50	45	44	42	16	50	50	42

Note:  $P$  denotes the number of free parameters in the true model.

- the Programme of International Assessment of Adult Competences (PIAAC)
  - **the problem solving** in technology-based environments domain
  - items are interactive and require a constructed response (**no guessing parameter**)
  - two computer-based problem solving modules each consisting of 7 items (7 + 7 intotal)
  - the problem solving modules + a module from a different domain
  - both problem solving modules  
(overall time limit of 30 minutes / module)
  - **data files:** 12th of June 2018, Canada (the largest number of respondents, 10315)
- the RA scores
  - the items were coded as correct/incorrect
  - eigenvalues of the correlation matrix: **one dimension** should be sufficient

1. the CI-HM ( $\mathcal{M}_1, \eta_0 = \eta_1$ )

- Gibbs Sampler with 20,000 iterations  
(including 10,000 burn-in, and a thinning of 2 was applied)

whether the model adequately captured the differences between the RTs of correct and incorrect responses?

- Posterior predictive checks (100 replicated data sets)

D1: differences between  $M_{\text{correct log-RT}}$  and  $M_{\text{incorrect log-RT}}$

D2: the ratio between  $S^2_{\text{correct log-RT}}$  and  $S^2_{\text{incorrect log-RT}}$

D3: the ratio between the first eigenvalues of the correlation matrices of log-RTs computed separately for correct and incorrect responses

## 1. Results

– the observed ones,

$$D1 = 0.338$$

$$D2 = 0.537$$

$$D3 = 1.285$$

- ➡ in all of the 100 generated data sets:
- ✓ D1 & D3: smaller than the observed ones
  - ✓ D2: larger than the observed one

➡ there is likely CD between RA and RT



## 2. fitted two CD models

– the first model:

$$f(x_i^*, t_i^* | \theta, \eta) = \mathcal{N}_2 \left( \begin{bmatrix} \alpha_i \theta + \beta_i \\ \xi_i - \lambda_i \eta \end{bmatrix}, \begin{bmatrix} 1 & \rho_i \sigma_i \\ \rho_i \sigma_i & \sigma_i^2 \end{bmatrix} \right)$$

– the second model:

$$f(x_i, t_i | \boldsymbol{\theta}, \boldsymbol{\eta}) = f(t_i | \boldsymbol{\theta}, \boldsymbol{\eta}) f(x_i | t_i, \boldsymbol{\theta}, \boldsymbol{\eta})$$

– Gibbs Samplers:

(20,000 iterations, including 10,000 burn-in, and a thinning of 2)

## 2. Results

Model	$P$	mAIC	$\bar{D}_1$	$\bar{D}_2$	$\bar{D}_3$	$p_1$	$p_2$	$p_3$
CI model ( $\mathcal{M}_1, \eta_0 = \eta_1$ )	71	234,633.3	0.215	0.779	0.811	.00	1.00	.00
CD model from approach 1	85	229,347.3	0.323	0.887	0.847	.00	1.00	.00
CD model from approach 2	99	228,351.8	0.322	0.819	0.837	.01	1.00	.00

3. fitted the set of two-dim models and the set of three-dim models
  - Gibbs Samplers:  
(20,000 iterations, including 10,000 burn-in, and a thinning of 2)

## 3. Results

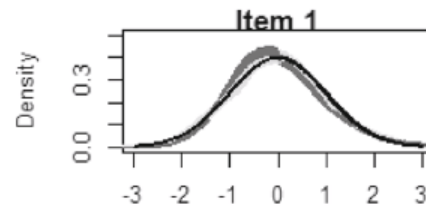
➔ the observed ones:  $D_1 = 0.338$ ,  $D_2 = 0.537$ ,  $D_3 = 1.285$

Model	$P$	mAIC	$\bar{D}_1$	$\bar{D}_2$	$\bar{D}_3$	$p_1$	$p_2$	$p_3$
CI model ( $\mathcal{M}_1, \eta_0 = \eta_1$ )	71	234,633.3	0.215	0.779	0.811	.00	1.00	.00
CD model from approach 1	85	229,347.3	0.323	0.887	0.847	.00	1.00	.00
CD model from approach 2	99	228,351.8	0.322	0.819	0.837	.01	1.00	.00
CD models from approach 3								
$\mathcal{M}_2, \eta_0 = \eta_1$	85	227,699.7	0.336	0.859	0.832	.36	1.00	.00
$\mathcal{M}_{3a}, \eta_0 = \eta_1$	99	225,809.6	0.335	0.707	0.650	.32	1.00	.00
$\mathcal{M}_{4a}, \eta_0 = \eta_1$	113	218,993.7	0.336	0.549	1.240	.34	.94	.08
$\mathcal{M}_1, \eta_0 \neq \eta_1$	75	229,599.2	0.377	0.652	0.663	1.00	1.00	.00
$\mathcal{M}_2, \eta_0 \neq \eta_1$	88	225,194.5	0.336	0.733	0.667	.34	1.00	.00
$\mathcal{M}_{3a}, \eta_0 \neq \eta_1$	101	224,114.3	0.336	0.693	0.796	.29	1.00	0.00
$\mathcal{M}_{4a}, \eta_0 \neq \eta_1$	105	216,824.9	0.337	0.541	1.250	.40	.64	.15

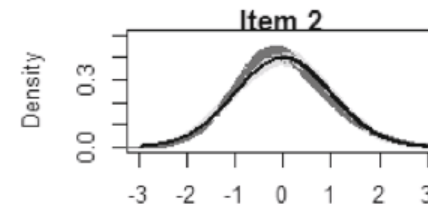
➔ PPC (100 data sets):  $p$ -value of 1

- assumes a separate lognormal distribution for RT for the two RA outcomes
- examined the posterior distribution of the standardized residuals of log-RTs

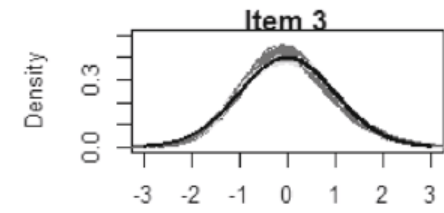
- three-dimensional  $\mathcal{M}_{4a}$



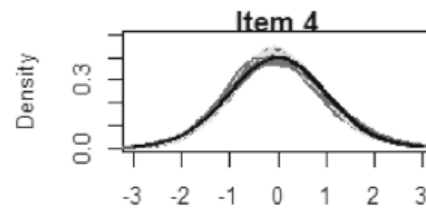
Standardised residual log-RT



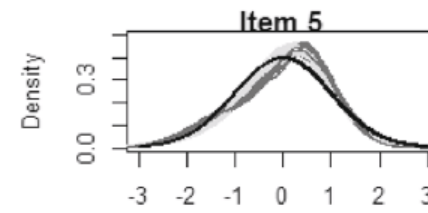
Standardised residual log-RT



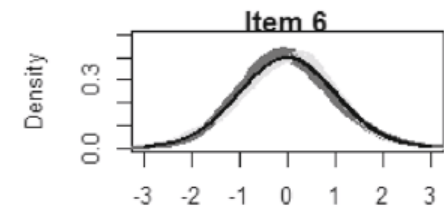
Standardised residual log-RT



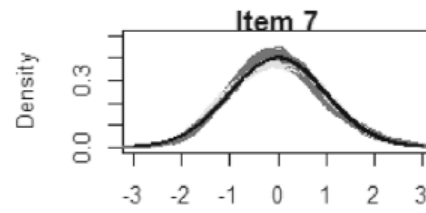
Standardised residual log-RT



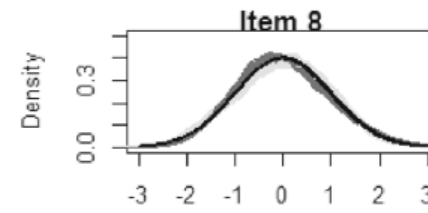
Standardised residual log-RT



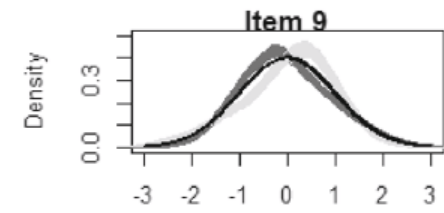
Standardised residual log-RT



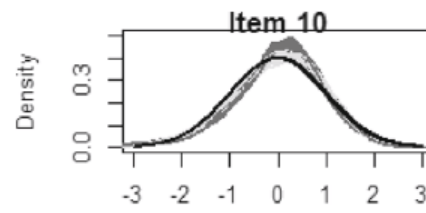
Standardised residual log-RT



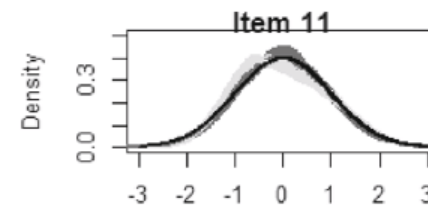
Standardised residual log-RT



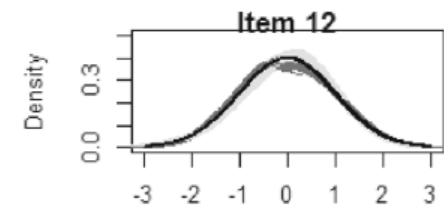
Standardised residual log-RT



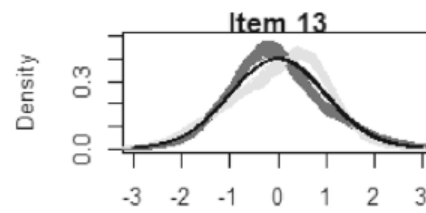
Standardised residual log-RT



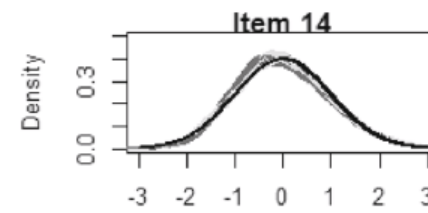
Standardised residual log-RT



Standardised residual log-RT

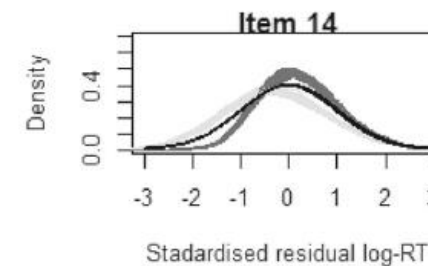
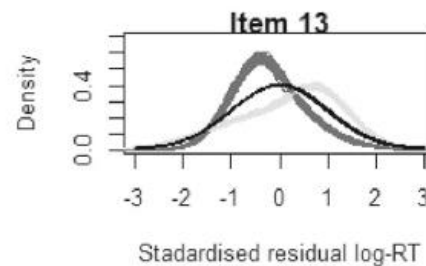
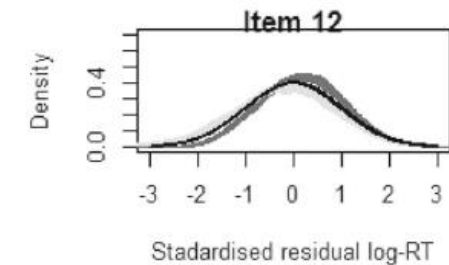
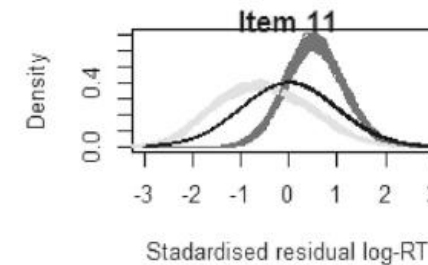
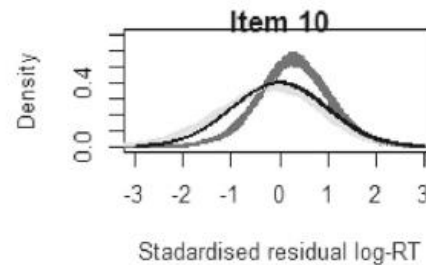
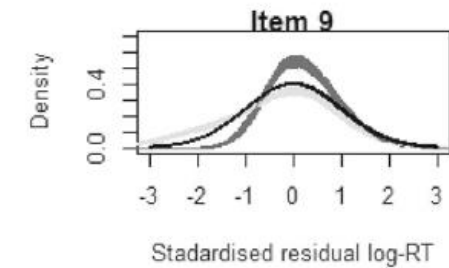
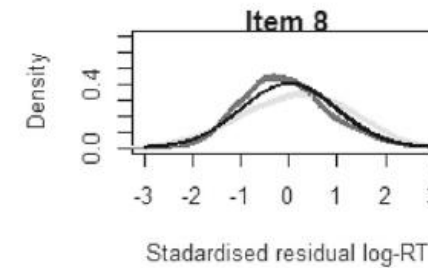
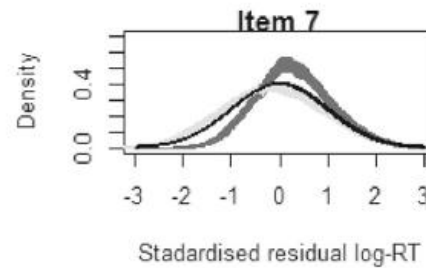
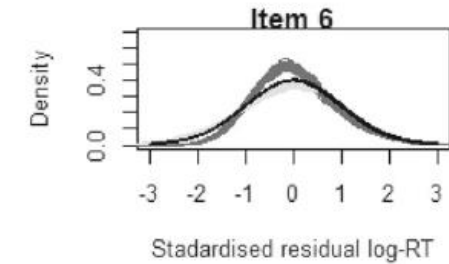
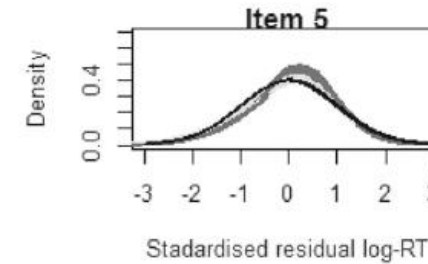
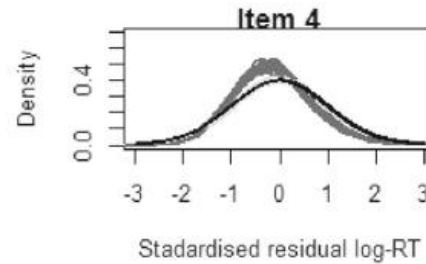
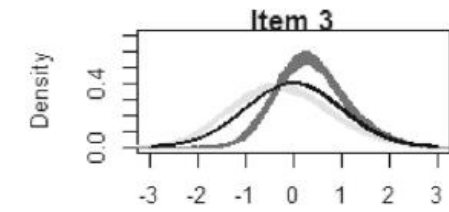
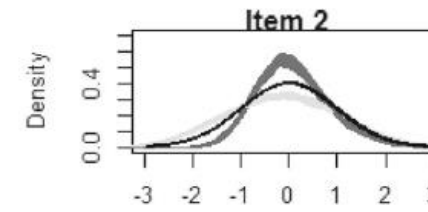
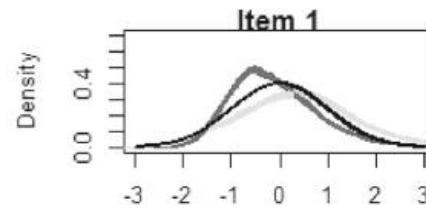


Standardised residual log-RT



Standardised residual log-RT

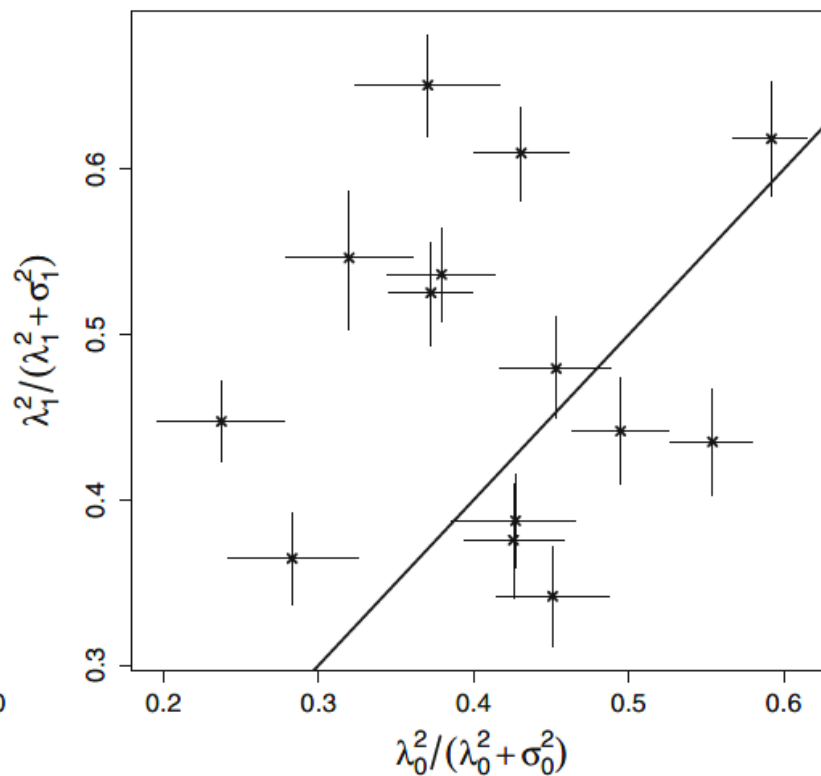
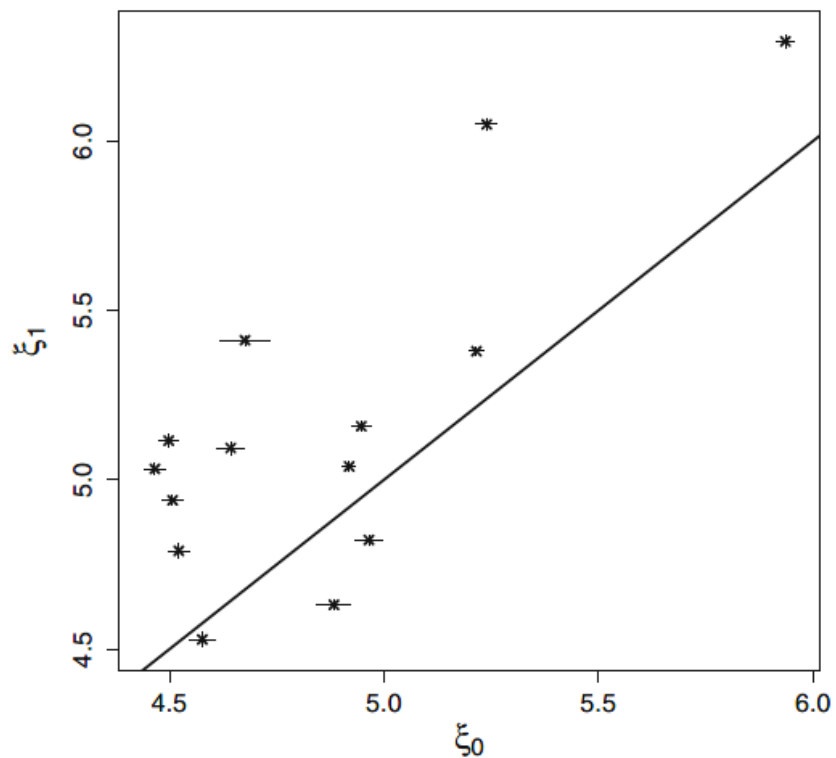
- the CI model



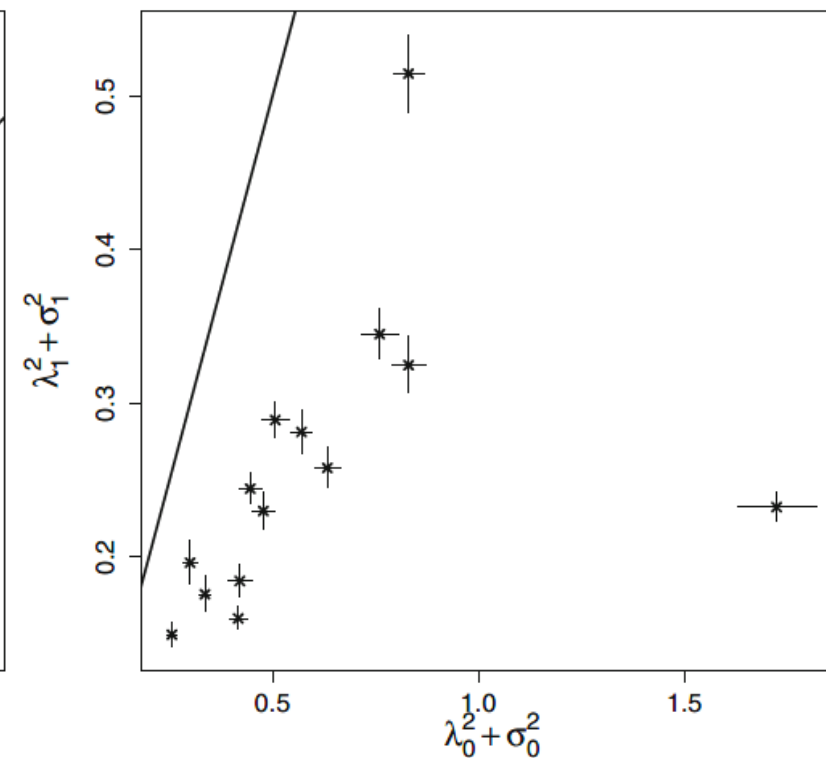
- three correlations between the person parameters
  - $\theta$  and  $\eta_0$ : -0.661 [-.642,-.679]
    - persons who give fast incorrect responses generally having a lower ability level
  - $\theta$  and  $\eta_1$ : 0.038 [.005,.072]
    - response speed and ability is much weaker
  - $\eta_0$  and  $\eta_1$ : 0.689 [.662,.714]
    - the two speed latent variables are strongly associated but still only share less than 50% of their variance



- the 95% credible intervals for relevant item properties



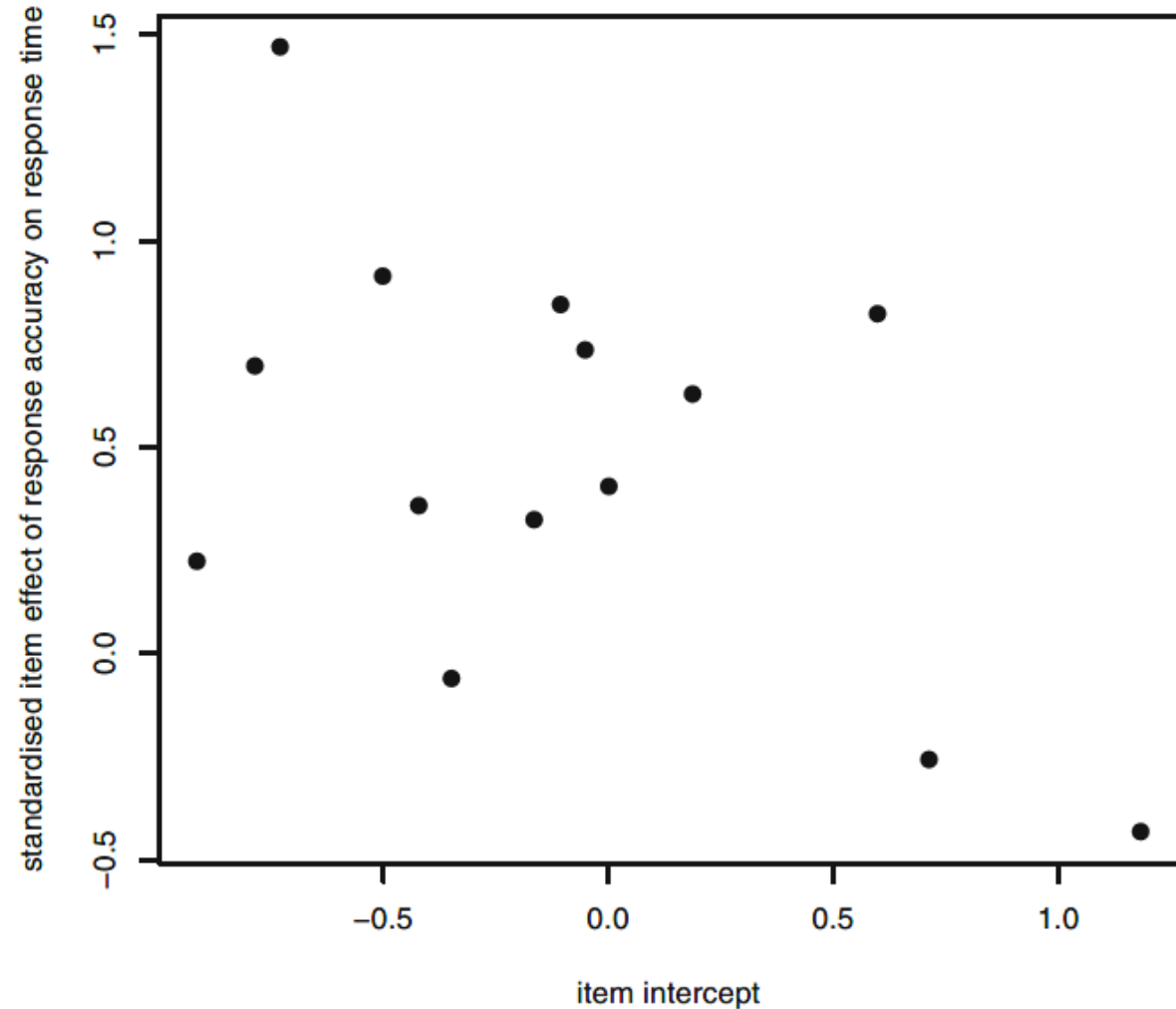
variance explained by the speed



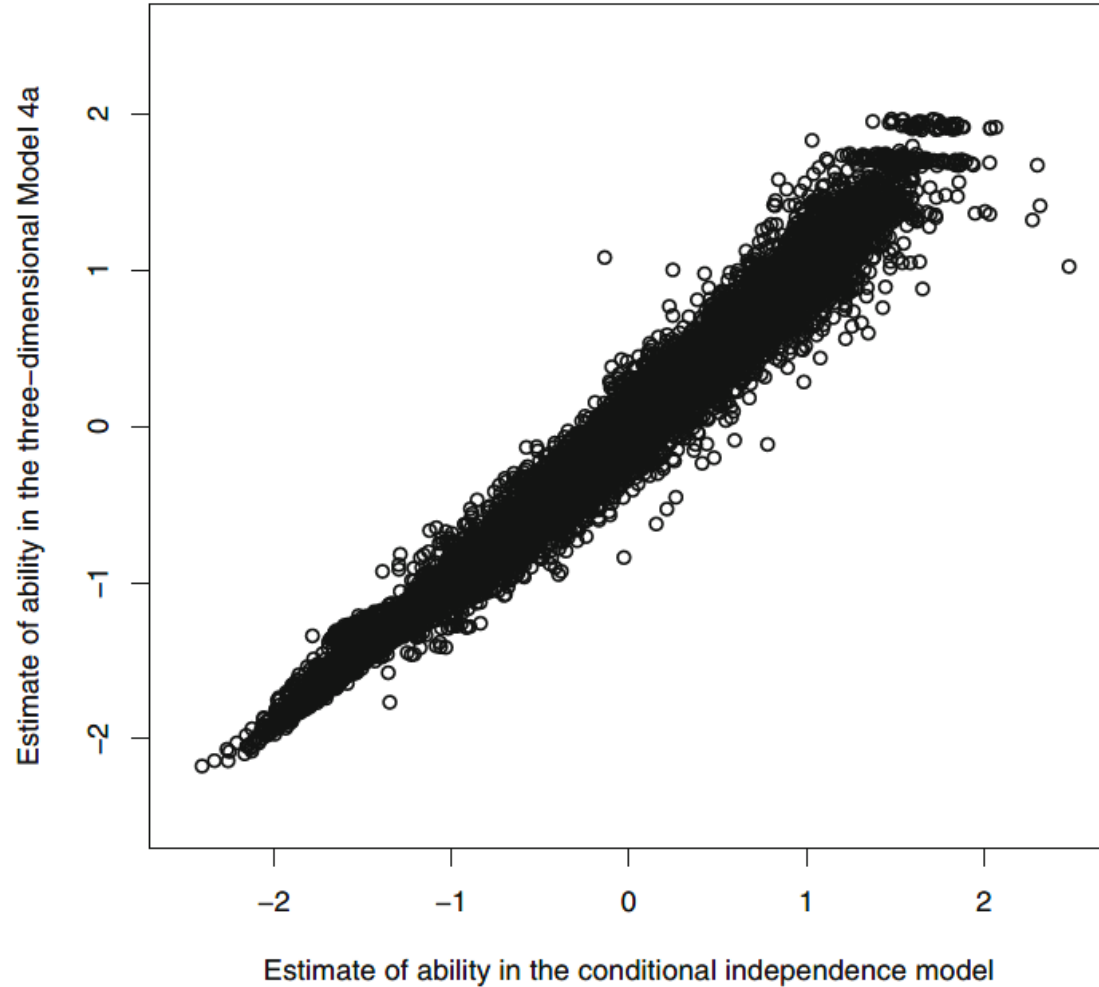
total variance

- quantify the strength and direction of the CD
  - item-specific standardized effect of RA on log-RT

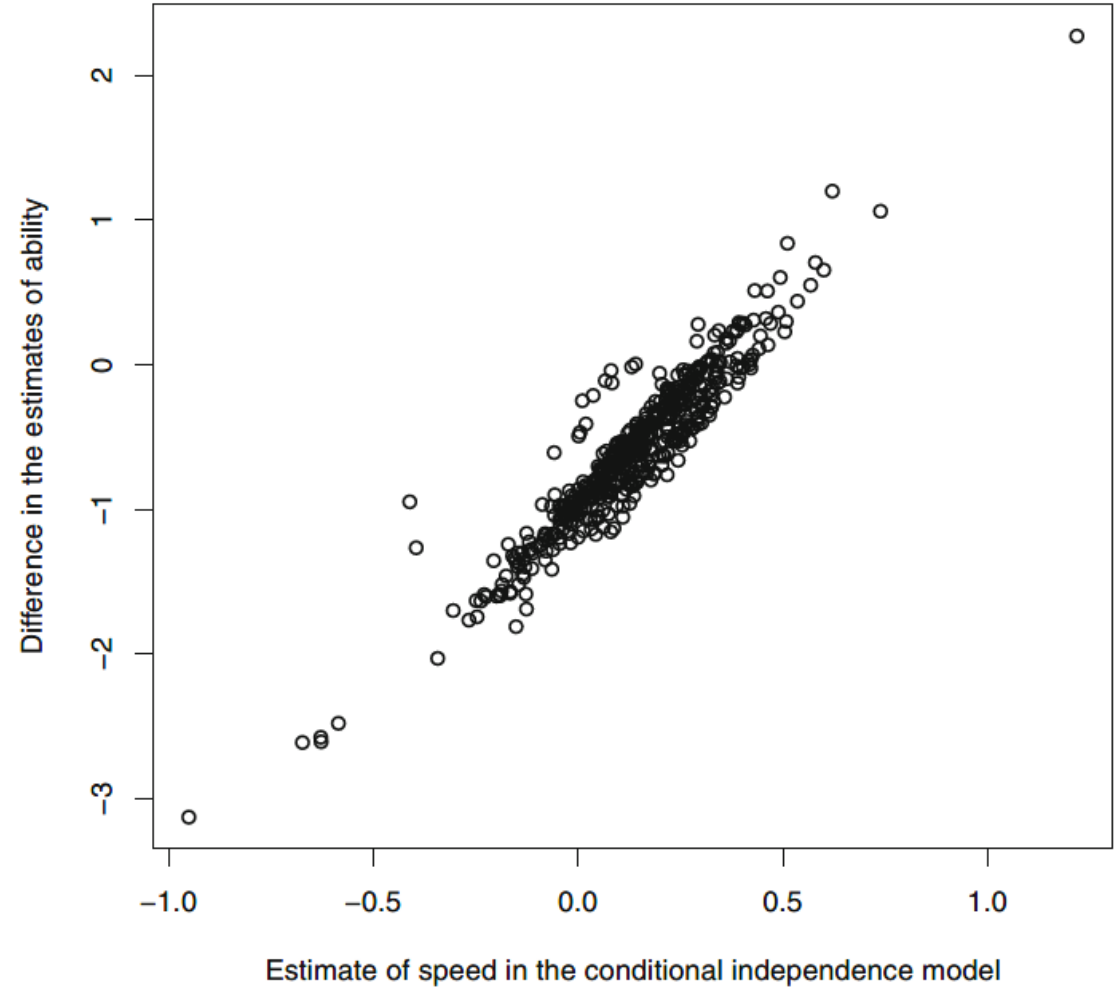
$$\frac{\xi_{i1} - \xi_{i0}}{\sqrt{\frac{(N_{i0} - 1)(\sigma_{i0}^2 + \lambda_{i0}^2) + (N_{i1} - 1)(\sigma_{i1}^2 + \lambda_{i1}^2)}{N_{i0} + N_{i1} - 2}}}}$$



All persons



Persons with all correct



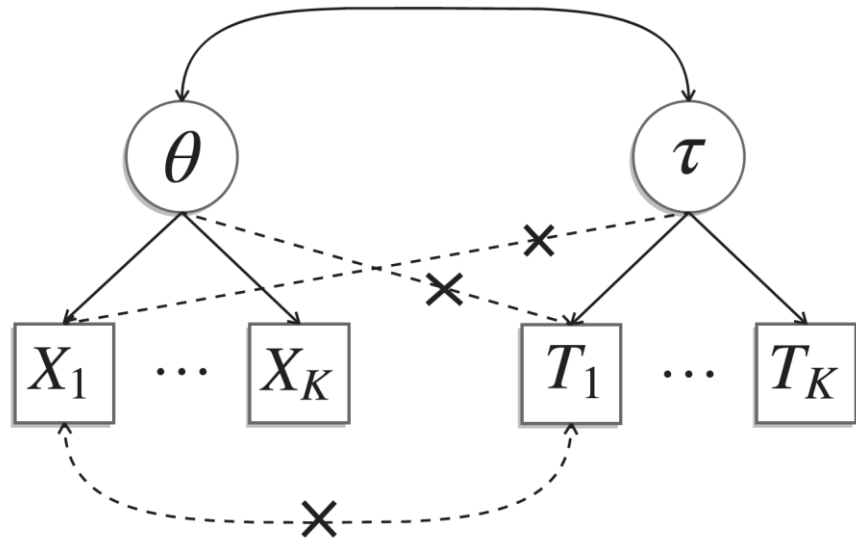
- sample sizes:
  - $N = 500$ ,  $N = 1000$ , and  $N = 2000$
  - a condition with the same sample size ( $N = 10,245$ ) and the pattern of missingness
- data generation:
  - The RA: the 2PNO model
  - The RT: the three-dimensional  $\mathcal{M}_{4a}$
- estimation:
  - Gibbs Sampler  
(6000 iterations, including 1000 burn-in, and a thinning of 2)

Condition	$\alpha$	$\beta$	$\xi$	$\lambda$	$\sigma^2$	$\Sigma_{12}$	$\Sigma_{13}$	$\Sigma_{23}$
Bias								
$N = 500$	0.033	0.011	0.004	0.001	0.002	0.016	0.019	0.024
$N = 1000$	0.020	0.006	0.001	0.001	0.001	0.011	0.011	0.016
$N = 2000$	0.009	0.003	0.002	0.001	0.001	0.011	0.004	0.009
$N = 10,245$	0.004	0.001	0.001	0.001	0.000	0.005	0.005	0.006
Variance								
$N = 500$	0.019	0.010	0.002	0.002	0.001	0.001	0.004	0.002
$N = 1000$	0.009	0.005	0.001	0.001	0.001	0.001	0.002	0.001
$N = 2000$	0.004	0.002	0.000	0.000	0.000	0.000	0.001	0.000
$N = 10,245$	0.002	0.001	0.000	0.000	0.000	0.000	0.000	0.000
Mean squared error								
$N = 500$	0.021	0.010	0.002	0.002	0.001	0.001	0.004	0.002
$N = 1000$	0.010	0.005	0.001	0.001	0.001	0.001	0.002	0.001
$N = 2000$	0.004	0.002	0.000	0.000	0.000	0.000	0.001	0.001
$N = 10,245$	0.002	0.001	0.000	0.000	0.000	0.000	0.000	0.000

Note: The first three conditions have a complete design, and the last condition has an incomplete design with the missingness patterns matching those in the empirical example.

- proposed a framework to directly investigate the differences of RTs between correct and incorrect responses
- all model parameters can generally be **recovered well** if the model is correctly specified
- **the mAIC** with a posterior predictive check is well-suited for selecting the correct model
- there may in practice be **notable relevant differences** between the models for the RTs of correct and incorrect responses
- **two speed** latent variables were needed to best model the empirical data

- **other parametric forms** for the RT model for correct and incorrect responses could be explored



ability latent variable(s) is correctly specified,  
**confounding**

make it possible to deal with **polytomously**

- it is still assumed that RTs only provide collateral information for the estimation of ability through the speed latent variable(s) in the model

THANKS FOR LISTENING!

REPORTER

YINGSHI HUANG