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#### MAPPING UNOBSERVED ITEM-RESPONDENT INTERACTIONS: A LATENT SPACE ITEM RESPONSE MODEL WITH INTERACTION MAP



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- Item response theory (IRT)
	- $-$  respondent  $j$  & item  $i$

 $logit(\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i)) = \alpha_j + \beta_i$ 

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- 2. consistency of success probability (same ability, same easiness)

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unobserved interaction / dependence



(same ability, different cultures)

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unobserved interaction / dependence

- Purpose:
	- − introduce a novel latent space model



# **Latent Space Model** 6

• How to interpret dependence?

*i*1



Items answered by examinee  $j$ 

# **Latent Space Model**

• How to interpret dependence?



Responses of different examinees on item i

# **Latent Space Model** 8

• How to interpret dependence: under a network structure



#### dependence

• Capture local item and person dependence: doubly LSM



Items answered by examinee *j* 



Responses of different examinees on item  $i$ 

• Capture local item and person dependence: doubly LSM



Items answered by examinee *j* 

#### **item pairs**



examinee  $j$  who answers item 2, 3, and 6 correctly

 $U_{j,I\times I} = \{x_{ji}x_{jk}\}\$  (item *i* and item *k*)

• Capture local item and person dependence: doubly LSM



Responses of different examinees on item i

**examinee pairs**



item  $i$  being answered correctly only by examinee 1, 2, and 9

 $\boldsymbol{Y}_{i,N\times N}=\left\{x_{ji}x_{li}\right\}$  (examinee  $j$  and examinee  $l$ )

• How to model the probability of a relation between nodes: embedding

#### the concept of "social space"

*a space of unobserved latent characteristics that represent potential transitive tendencies in network relations*

• How to model the probability of a relation between nodes: embedding

#### the concept of "social space"

*a space of unobserved latent characteristics that represent potential transitive tendencies in network relations*

− an (maybe unappropriated) illustration:



These students came from the same school

(social space: school)



*Ick Hoon Jin, Minjeong Jeon, 2019 Psychometrika Peter D. Hoff, Adrian E. Raftery, Mark S. Handcock, 2002 JASA*

• *p*-dimensional latent space (typically 2-3 dimensions)



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• A concise version is needed



- 1. work with *functions* of item response data
- 2. deal with *multiple* networks
- 3. must *combine* two LSMs for simultaneous estimation:  $w_j = f_j(Z)$

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- 1. work with *functions* of item response data
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- 3. must *combine* two LSMs for simultaneous estimation:  $w_i = f_i(\mathbf{Z})$

work with original response data & single network

- Work with original response data & single network
	- − original response data:  $Y \in \{0, 1\}^{N \times I}$
	- − a single network?



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- Work with original response data & single network
	- − original response data:  $Y \in \{0, 1\}^{N \times I}$
	- − a single network:  $logit(\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i, a_j, b_i)) = \alpha_j + \beta_i + g(a_j, b_i)$

the position of respondent  $i$  and item  $i$ 

- Work with original response data & single network
	- − original response data:  $Y \in \{0, 1\}^{N \times I}$
	- − a single network:  $logit(\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i, a_j, b_i)) = \alpha_j + \beta_i + g(a_j, b_i)$

the position of respondent  $i$  and item  $i$ 

1. multiplicative effect 
$$
g(a_j, b_i) = a_j^{\top} b_i
$$

2. distance effect 
$$
g(a_j, b_i) = -\gamma d(a_j, b_i)
$$
  
weight distance  
 $(\ge 0)$ 

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$$

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- 1. multiplicative effect  $g(a_j, b_i) = a_j^{\top} b_i$
- 2. distance effect  $g(a_j, b_i) = -\gamma d(a_j, b_i)$ weight distance  $( \geq 0)$

- reflexivity:  $d(a, b) = 0$  if and only if  $a = b \in M$ ;
- symmetry:  $d(a, b) = d(b, a)$  for all  $a, b \in M$ ;
- triangle inequality:  $d(a, b) \leq d(a, c) + d(b, c)$  for all  $a, b, c \in M$ .

a) 
$$
\ell_1
$$
-distance:  $d(a, b) = ||a - b||_1 = \sum_{i=1}^p |a_i - b_i|$ 

b) 
$$
\ell_2
$$
-distance:  $d(a, b) = ||a - b||_2 = \sqrt{\sum_{i=1}^p (a_i - b_i)^2}$ 

c)  $\ell_{\infty}$ -distance:  $d(a, b) = ||a - b||_{\infty} = \max_{1 \leq i \leq p} |a_i - b_i|$ 

- Work with original response data & single network
	- − original response data:  $Y \in \{0, 1\}^{N \times I}$
	- − a single network:

$$
logit(\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i, \boldsymbol{a}_j, \boldsymbol{b}_i)) = \alpha_j + \beta_i + g(\boldsymbol{a}_j, \boldsymbol{b}_i)
$$

the position of respondent  $i$  and item  $i$ 

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$$
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2. distance effect 
$$
g(a_j, b_i) = -\gamma d(a_j, b_i)
$$
  
weight distance  
 $(\ge 0)$ 

 $\gamma = 0$ : Rasch model  $y > 0$ : capture deviations from the main effects (embed respondents into a shared metric space)

• The distance effect is easier to interpret than the multiplicative effect

$$
a_j = (0, 1/100)
$$
  
\n
$$
b_i = (1/100, 0)
$$
  
\n
$$
a_j = (0, 100)
$$
  
\n
$$
b_i = (100, 0)
$$

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$$
  
\n
$$
a_j = (0, 100)
$$
  
\n
$$
b_i = (100, 0)
$$

- 1. multiplicative effect  $\mathbf{a}_i^{\top} \mathbf{b}_i = 0$  $\mathbf{a}_i^{\top} \mathbf{b}_i = 0$
- $d(a_j, b_i) = 141.42$ 2. distance effect  $d(a_j, b_i) = 0.01$  $(\ell_2$ -distance)

angle & length!

- Two main lines of research involving networks
	- 1. Graphical models (model structure):

vertices: the variables of interest

edge: conditional dependencies among variables

2. Random graph models (data structure):

vertices: individuals

edge: friendships among individuals / correct responses

- Practical advantages
	- − provides a geometric representation of interactions

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	- − provides a geometric representation of interactions

 $logit(\mathbb{P}(y_{i,i} = 1 | \alpha_i, \beta_i, \zeta_{i,i})) = \alpha_i + \beta_i + \zeta_{i,i}$ 

1. Rasch:

14 items & 200 respondents  $\zeta_{i,i} = 0$ 

2. Rasch with local dependence:

first 100 respondents  $\rightarrow$  item 1-7  $\zeta_{i,i} \sim N(2, 0.2)$  item 8-14  $\zeta_{i,i} = 0$ last 100 respondents  $\rightarrow$  item 8-14  $\zeta_{i,i} \sim N(2, 0.2)$  item 1-7  $\zeta_{i,i} = 0$ 

latent space dimension:  $M = \mathbb{R}^2$ 

- Practical advantages
	- − provides a geometric representation of interactions



- Theoretical advantages
	- − weaker conditional independence assumption

$$
\mathbb{P}(Y = y \mid \boldsymbol{\alpha}, \boldsymbol{\beta}, \gamma, A, B) = \prod_{j=1}^N \prod_{i=1}^I \mathbb{P}(Y_{j,i} = y_{j,i} \mid \alpha_j, \beta_i, \gamma, a_j, b_i)
$$

$$
\boldsymbol{\alpha} = (\alpha_1, \ldots, \alpha_N), \boldsymbol{\beta} = (\beta_1, \ldots, \beta_I), \boldsymbol{A} = (\boldsymbol{a}_1, \ldots, \boldsymbol{a}_N), \text{and } \boldsymbol{B} = (\boldsymbol{b}_1, \ldots, \boldsymbol{b}_I)
$$

allow for respondent–item interactions:

- $\checkmark$  testlets (e.g., items similar in content)
- $\checkmark$  learning and practice effects
- $\checkmark$  repeated measurements
- $\checkmark$  nested respondents (shared school or family memberships)

- Theoretical advantages
	- − drop some of the homogeneity assumptions same abilities: respondents  $j_1$  and  $j_2$

different distances from item *i*:  $d(a_{j_1}, b_i) < d(a_{j_2}, b_i)$ 

- Theoretical advantages
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different distances from item *i*:  $d(a_{j_1}, b_i) < d(a_{j_2}, b_i)$ 



- Theoretical advantages
	- − drop some of the homogeneity assumptions



(a) Two response patterns

Respondents 51-100 give correct responses to Items 4-6 only

Respondents 1–50 give correct responses to Items 1–3 only

- Theoretical advantages
	- − drop some of the homogeneity assumptions



(b) Two response patterns, with randomness



#### **Related Models** 36

- Other models with relaxed assumptions
	- − polytomous item models
	- − testlet and bifactor models
	- − finite mixture models

− …

- 1. require knowledge of the interaction structure
- 2. heterogeneity between latent classes, but assume homogeneity within latent classes
# **Related Models** 37

- Other models with interactions among respondents and items
	- − two-parameter IRT model

 $logit(\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i, \lambda_i)) = \lambda_i \alpha_j + \beta_i$ 

cannot visualize interactions

# **Related Models** 38

- Other models with interactions among respondents and items
	- − two-parameter IRT model

```
logit(\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i, \lambda_i)) = \lambda_i \alpha_j + \beta_i
```
cannot visualize interactions

− interaction IRT model

$$
logit(\mathbb{P}(Y_{j,i} = 1 \mid \alpha_j, \beta_i, \epsilon_{j,i})) = \alpha_j + \beta_i + \epsilon_{j,i}
$$

latent space model can be viewed as a special case:  $\epsilon_{j,i} = -\gamma d(a_j, b_i)$ 

the restrictions 1. facilitate estimation 2. make sense in practice (capture transitivity)

# **Related Models** 39

- Other models with interactions among respondents and items
	- − bilinear mixed effects models and related models

the multiplicative effects version

− differential item functioning

an interaction term is formed with a known categorical attribute of respondents (e.g., gender) and an item indicator



# **Bayesian Inference**

• Markov Chain Monte Carlo (MCMC)

logit( $\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i, \mathbf{a}_j, \mathbf{b}_i) = \alpha_j + \beta_i + g(\mathbf{a}_j, \mathbf{b}_i)$  $g(a_j, b_i) = -\gamma d(a_j, b_i)$ 

− (Posterior)∼ (Prior) (Likelihood)

# **Bayesian Inference** 41

• Markov Chain Monte Carlo (MCMC)

$$
logit(\mathbb{P}(Y_{j,i} = 1 \mid \alpha_j, \beta_i, a_j, b_i)) = \alpha_j + \beta_i + g(a_j, b_i)
$$
  

$$
g(a_j, b_i) = -\gamma \ d(a_j, b_i)
$$

− (Posterior)∼ (Prior) (Likelihood)

$$
f(\boldsymbol{\alpha}, \boldsymbol{\beta}, \boldsymbol{\gamma}, \boldsymbol{A}, \boldsymbol{B} | \boldsymbol{y}) \propto \left[ \prod_{j=1}^{N} f(\alpha_j) \right] \left[ \prod_{i=1}^{I} f(\beta_i) \right] f(\boldsymbol{\gamma}) \left[ \prod_{j=1}^{N} f(a_j) \right] \left[ \prod_{i=1}^{I} f(b_i) \right]
$$

$$
\times \left[ \prod_{j=1}^{N} \prod_{i=1}^{I} \mathbb{P} \left( Y_{j,i} = y_{j,i} | \alpha_j, \beta_i, \gamma, a_j, b_i \right) \right],
$$

# **Bayesian Inference**

• Markov Chain Monte Carlo (MCMC)

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− (Posterior)∼ (Prior) (Likelihood)

#### priors:

$$
\alpha_j \mid \sigma^2 \qquad \stackrel{\text{ind}}{\sim} \text{N}(0, \sigma^2), \quad \sigma^2 > 0, \quad j = 1, \dots, N
$$
\n
$$
\beta_i \mid \tau_\beta^2 \qquad \stackrel{\text{ind}}{\sim} \text{N}(0, \tau_\beta^2), \quad \tau_\beta^2 > 0, \quad i = 1, \dots, I
$$
\n
$$
\log \gamma \mid \mu_\gamma, \tau_\gamma^2 \sim \text{N}\left(\mu_\gamma, \tau_\gamma^2\right), \quad \mu_\gamma \in \mathbb{R}, \quad \tau_\gamma^2 > 0
$$
\n
$$
\sigma^2 \mid a_\sigma, b_\sigma \qquad \sim \text{Inv-Gamma}(a_\sigma, b_\sigma), \quad a_\sigma > 0, \quad b_\sigma > 0
$$
\n
$$
a_j \qquad \stackrel{\text{iid}}{\sim} \text{MVN}_p(0, I_p), \quad j = 1, \dots, N
$$
\n
$$
b_i \qquad \stackrel{\text{iid}}{\sim} \text{MVN}_p(0, I_p), \quad i = 1, \dots, I,
$$

• Identifiability

a) 
$$
\ell_1
$$
-distance:  $d(a, b) = ||a - b||_1 = \sum_{i=1}^p |a_i - b_i|$   
b)  $\ell_2$ -distance:  $d(a, b) = ||a - b||_2 = \sqrt{\sum_{i=1}^p (a_i - b_i)^2}$   
c)  $\ell_{\infty}$ -distance:  $d(a, b) = ||a - b||_{\infty} = \max_{1 \le i \le p} |a_i - b_i|$ 

undirected: distances are inherently symmetric

invariant to translations, reflections, and rotations

- Identifiability
	- − post-processing the MCMC output with Procrustes matching



- Identifiability
	- − post-processing the MCMC output with Procrustes matching



observed matrix target matrix  $B$ 

find a transformation  $T$ , to produce greatest similarity between  $AT$  and  $B$ 



relative distances between positions

three transformation steps of an ordinary Procrustes fit for two configurations (from WiKi)

#### • Model selection

 $logit(\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i, a_j, b_i)) = \alpha_j + \beta_i + g(a_j, b_i)$  $g(a_j, b_i) = -\gamma d(a_j, b_i)$ 

 $\gamma = 0$  or  $\gamma > 0$  ?

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 $\gamma = 0$  or  $\gamma > 0$  ?

− spike-and-slab prior

 $\gamma$  is likely to be sampled from N(0,0.01) or N(0,1)?



*Fig made by Fabian Dablander*

#### • Model selection

 $logit(\mathbb{P}(Y_{j,i} = 1 | \alpha_j, \beta_i, a_j, b_i)) = \alpha_j + \beta_i + g(a_j, b_i)$  $g(a_i, b_i) = -\gamma d(a_i, b_i)$ 

 $\gamma = 0$  or  $\gamma > 0$  ?





# **Simulation Study**

- The accuracy of model selection approach
	- 1. Rasch:

14 items & 200 respondents

100 datasets with  $\gamma = 0$ 

2. latent space model:

first 100 respondents  $\rightarrow$  item 1-7 last 100 respondents  $\rightarrow$  item 8-14 100 datasets with  $\gamma = 1.7$ 

 $\sim$  compute the proportion of times  $\delta = 1$  in the MCMC posterior sample

## **Simulation Study**

(a) Truth:  $\gamma = 0$ (b) Truth:  $\gamma = 1.7$  $300 60 200 40 100 20 0 0 -$ 0.985  $0.00$  $0.05$  $0.10$  $0.15$ 0.990 0.995 Posterior inclusion probability Posterior inclusion probability

Histogram of the estimated posterior probability of the event  $\delta = 1$ , called "posterior inclusion probability." a Data are generated from the Rasch model with  $\gamma = 0$ . **b** Data are generated from the latent space model with  $\gamma = 1.7$ 

#### • Data and estimation

- 1. The woman decides on her own that she does not wish to have the child
- 2. The couple agree that they do not wish to have the child
- 3. The woman is not married and does not wish to marry the man
- 4. The couple cannot afford any more children.
- 5. There is a strong chance of a defect in the baby
- 6. The woman's health is seriously endangered by the pregnancy
- 7. The woman became pregnant as a result of rape

0.42, 0.52, 0.47, 0.53, 0.86, 0.94, 0.93

− 20,000 iterations (10,000 burn-in)

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0.42, 0.52, 0.47, 0.53, 0.86, 0.94, 0.93

- − 20,000 iterations (10,000 burn-in)
- − Convergence: the scale reduction factor < 1.06
- $-$  Model selection: probability of  $\delta = 1$  was 0.99



### **Example 1: Attitudes to Abortion** 53





ID	$\mathbf{I}$	13	I4	15	16	
27						
92						
132						
191						
273						
330						
653						
662						
675						

the region of  $X > 0.3$  and  $Y > 0.025$ 

These people tend to give positive responses to I1–I4, but negative responses to I5–I7

(a) LS  $\beta_i$ (b) Rasch  $\beta_i$  $\overline{ }$ ဖာ ဖ  $\blacktriangleleft$ ဖာ က  $\rightarrow$  $\sim$  ო –  $\overline{ }$  $\sim$  - $\circ$  -Y.  $\overline{\phantom{0}}$  $\overline{\mathbf{a}}$  $b<sub>2</sub>$  $\overline{3}$  $\frac{1}{4}$ aq бő  $\overline{7}$  $b7$  $b<sub>1</sub>$  $b2$  $b<sub>3</sub>$  $b<sub>4</sub>$  $b<sub>5</sub>$  $b6$ 

# **Example 1: Attitudes to Abortion** 56



## **Example 1: Attitudes to Abortion** 57



Predicted proportions of the positive responses for the seven items for the attitudes to abortion data. The red dot in each box indicates the proportion of positive responses calculated from the raw data

- Data and estimation
	- − the Competence Profile Test of Deductive Reasoning—Verbal assessment (DRV)
	- − 24 binary items (0 = correct, 1 = incorrect)
	- − 418 school students (162 female)



- Data and estimation
	- − the Competence Profile Test of Deductive Reasoning—Verbal assessment (DRV)
	- − 24 binary items (0 = correct, 1 = incorrect)
	- − 418 school students (162 female)

- − 20,000 iterations (10,000 burn-in)
- − Convergence: the scale reduction factor < 1.1
- $-$  Model selection: probability of  $\delta = 1$  was 0.99

move forward with the latent space model



a 95% posterior credible intervals for the  $\beta_i$  estimates (b1 and b24 on the X-axis represent Items 1 to 24), and **b** the distribution of the  $\alpha_i$  estimates per total test score for the DRV data. The estimates are from the latent space model



Box plots of the predicted proportions of the correct responses for the 24 DRV test items from 10,000 replicated data. The red dot in each box indicates the proportion of the correct responses for the corresponding item from the raw data





$$
\gamma = 2.23 \; [2.08, 2.35]
$$

Members of the four item groups identified in the DRV data latent space



Numbers in parenthesis indicate item numbers. The acronyms in the item labels indicate the following design factors and their levels: (1) UN vs. N: no negation (UN) and Negation (N) for the presentation of the antecedent factor. (2) CO vs. AB vs. AC: Concrete (CO), Abstract (AB), and Counterfactual (CF) for the content of conditional factor. (3) MP vs. MT vs. NA vs. AC: Modus Ponens (MP), Modus Tollens (MT), Negation of Antecedent (NA), and Affirmation of Consequent (AC) for the type of inference factor

 $0.4$ 

 $0.3 -$ 

Deraity<br>0.2

 $0.1 -$ 

6.0

 $0.4$ 

 $0.3 -$ 

Density<br> $\frac{1}{2}$ 

 $0.1$ 

 $0.0 -$ 

o<br>Logit Prob



o<br>Logit Prob

− Cosine similarity between item groups

$$
\cos(\theta) = \frac{\mathbf{a}^\top \mathbf{b}}{||\mathbf{a}||_2 ||\mathbf{b}||_2}
$$

the angle between two vectors  $\boldsymbol{a}$  and  $\boldsymbol{b}$ 



(b) with item group vectors



Cosine similarity measures between (centers of) the four item groups



#### − Respondent structure









#### Type of Inference combined with **Concrete conditionals**

Rotated latent space for the DRV data with oblim rotation. Dots represent respondents and numbers represent items. Four item groups are distinguished with four different colors. I1: Items 2, 3, 6, 7; I2: Items 10, 11, 14, 15, 18, 19, 22, 23; I3: Items 9, 12, 13, 16, 17, 20, 21, 24; I4: Items 1, 4, 5,8 (Color figure online)

• interactions among respondents and items are present and non-negligible

- whether test items are differentiated or grouped together as blueprinted by test developers
	- (e.g., the Presentation of Antecedent barely contributed to item differentiation)
- detect unintended or undesirable forms of test-taking behavior (e.g., respondents that are located close to the last test items)
- provide feedback on the test performance (e.g., identify items that individual test takers may be struggling with)

# THANKS FOR ATTENTION!

REPORTER

YINGSHI HUANG